

# Math 525

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## 1 Simplest models

Suppose we have a finite  $\Omega : |\Omega| < \infty$ , all are equally likely.

$$\mathbb{P}(i) = \frac{1}{|\Omega|}.$$

$$E \in \Omega, \mathbb{P}(E) = \frac{|E|}{|\Omega|}.$$

Axioms: Easy! (just count the numbers.)

Here are some special cases:

Flip 2 coins, we have  $2^2 = 4$  outcomes, we can record the result as a vector  $(H/T, H/T)$ .

Here we can think of the sample space  $\Omega$  as a 2-dimensional table:

Notice that the top left and bottom right outcomes are different, and all four outcomes are equally likely.

Suppose  $E =$ "have at least one  $H$ ". We have

$$P(E) = \frac{3}{4}.$$

## 2 Sampling: Balls in urns

An urn with  $a$  balls red,  $b$  balls white.

Now we have to decide an experimental procedure:

- (i) with replacement
- (ii) without replacement.

Real life counterparts: sampling population of Michigan when you have a phone book.

Suppose  $A$  = "draw red",  $B$  = "draw white".

1st draw:

$$\mathbb{P}(A) = \frac{a}{a+b}.$$

2nd draw:  $\mathbb{P}(A)$  would be the same. Thus we have

$$\mathbb{P}(A_1 \cap A_2) = \left(\frac{a}{a+b}\right)^2.$$

This illustrates the idea of independence (formalized later.)

Now say we draw an "A" first (without replacement), then we have

$$\mathbb{P}(A_2 | A_1) = \frac{a-1}{a+b-1}$$

where  $\mathbb{P}(A_2 | A_1)$  is the conditional probability (the probability of  $A_2$  given that  $A_1$  happened.)

*Remark.* Finite probabilistic problems depend on counting and correctly interpreting the experimental protocol.

Recall the axioms:

$E \subset \Omega$ , then

1.  $\mathbb{P}(E) \in [0, 1]$
2.  $\mathbb{P}(\emptyset) = 0, \mathbb{P}(\Omega)$
3.  $\mathbb{P}(\bigcap_i A_i) = \sum_i \mathbb{P}(A_i)$ , where  $i$  is countable  $A_i \cap A_j = \emptyset$  for all  $i, j$ .

What about overlap?

$$\mathbb{P}(A \cup B) = \mathbb{P}(A) + \mathbb{P}(B) - \mathbb{P}(A \cap B).$$

Refer to Venn diagram.

Formula:

$$\mathbb{P}\left(\bigcup_i E_i\right) = \sum_i \mathbb{P}(E_i) - \sum_{i \neq j} \mathbb{P}(E_i \cap E_j) + \sum \mathbb{P}(E_i \cap E_j \cap E_k) - \dots$$

*Proof.* By induction. **Base case**  $n = 2$ .

$$A \cup B = (A \setminus (A \cap B)) \cup (B \setminus (A \cap B)) \cup (A \cap B)$$

Where all three sets are disjoint.

$$\begin{aligned}\mathbb{P}(A \cup B) &= \mathbb{P}(A \setminus (A \cap B)) + \mathbb{P}(B \setminus (A \cap B)) + \mathbb{P}(A \cap B) \\ &= \mathbb{P}(A) - \mathbb{P}(A \cap B) + \mathbb{P}(B) - \mathbb{P}(A \cap B) + \mathbb{P}(A \cap B) \\ &= \mathbb{P}(A) + \mathbb{P}(B) - \mathbb{P}(A \cap B).\end{aligned}$$

The rest follows by induct on  $n$ . (refer to textbook) ■

### 3 Independence & Conditional Probabilities

Consider 2 coin flips, all outcomes equal. Suppose  $\Omega$  is sample space,  $E =$ “at least one head”.

$$\mathbb{P}(2 \text{ Heads} \mid E) = \frac{1}{3}.$$

We can count this, but what is the formula?

$$\mathbb{P}(2 \text{ Heads} \mid E) = \frac{\mathbb{P}(2 \text{ Heads} \cap E)}{\mathbb{P}(E)} = \frac{\mathbb{P}(2 \text{ Heads})}{\mathbb{P}(E)} = \frac{1/4}{3/4} = \frac{1}{3}.$$

Clearly we have  $\mathbb{P}(E \mid E) = 1$ .

**Definition 3.1.** Conditional Probability.

$$\mathbb{P}(A \mid B) = \frac{\mathbb{P}(A \cap B)}{\mathbb{P}(B)}, (\mathbb{P}(B) \neq 0).$$

**Definition 3.2.** Independence of event  $A$  and  $B$ .

$$\mathbb{P}(A \mid B) = \mathbb{P}(A) \text{ i.e. } \mathbb{P}(A) = \frac{\mathbb{P}(A \cap B)}{\mathbb{P}(B)} \Leftrightarrow \mathbb{P}(A \cap B) = \mathbb{P}(A) \cdot \mathbb{P}(B).$$

*Remark.* Bayes Theorem.

$$\begin{aligned}\mathbb{P}(A \mid B) &= \frac{\mathbb{P}(A \cap B)}{\mathbb{P}(B)} \\ &= \frac{\mathbb{P}(B \cap A) \mathbb{P}(A)}{\mathbb{P}(B) \mathbb{P}(A)} \\ &= \frac{\mathbb{P}(A)}{\mathbb{P}(B)} \mathbb{P}(B \mid A).\end{aligned}$$

**Example 3.1.** “Compounds experiment” 3 coins,  $F : p = \frac{1}{2}, C : p = \frac{2}{3}, B : p = \frac{1}{3}$ . If  $F = H$  show  $C$ , otherwise show  $B$ .

Questions: what is the prob of flipping and seeing an H (=  $H_1$ )?

$$\begin{aligned}\mathbb{P}(H_1) &= \mathbb{P}(H_1 | F = H)\mathbb{P}(F = H) + \mathbb{P}(H_1 | F = T)\mathbb{P}(F = T) \quad (\text{partitioning into cases or conditioning.}) \\ &= \frac{2}{3} \cdot \frac{1}{2} + \frac{1}{3} \cdot \frac{1}{2} \\ &= \frac{1}{2}.\end{aligned}$$

What about  $H_1H_2$ ?